

Article

FEM-Based Compression Fracture Risk Assessment in Osteoporotic Lumbar Vertebra L1

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Abstract: This paper presents a finite element method (FEM)-based fracture risk assessment in patient-specific osteoporotic lumbar vertebra L1. The influence of osteoporosis is defined by variation of parameters such as thickness of the cortical shell, the bone volume–total volume ratio (BV/TV), and the trabecular bone score (TBS). The mechanical behaviour of bone is defined using the Ramberg–Osgood material model. This study involves the static and nonlinear dynamic calculations of von Mises stresses and follows statistical processing of the obtained results in order to develop the patient-specific vertebra reliability. In addition, different scenarios of parameters show that the reliability of the proposed model of human vertebra highly decreases with low levels of BV/TV and is critical due to the thinner cortical bone, suggesting high trauma risk by reason of osteoporosis.

Keywords: bone tissue; elastoplasticity; finite element method; fracture risk; osteoporosis; trabeculae; trabecular bone score; vertebra

1. Introduction

Spinal bones can be affected by several diseases, but spinal brittleness is mainly caused by osteoporosis. Due to the importance of studying this disease and its consequences, various social and medical aspects of osteoporosis have been investigated all over the world [1–4]. A comprehensive discussion of the role of computational biomechanical modeling is presented in a review article by Doblare et al. [5]. The paper emphasized numerical modeling as the main directions for future research. The adequacy of the modeling of finite elements strongly depends on the selection of the mechanical properties, the geometrical form of the numerical model, and mesh making possibilities.

With regard to considerable difficulties in the assignment of the anatomical geometry to the model, most researches are focused on simplifications. Some developments are restrained to small fragments of the vertebra [6,7], but most studies concerning the vertebral body in isolation [8–11] also use a simplified form, excluding the posterior elements [12,13].

A two-scale modeling approach was proposed in a series of papers where the macroscopic defining behavior is defined by the influence of the microstructure. For example, the work of McDonald et al. used a microscopic lattice network [12]. A similar approach was applied by Reference [14]. In contrast, a more comprehensive method of modeling was used in Reference [15] where the porous trabecular cell in the cancellous bone is considered as three-dimensional solid.

To determine the level of stress and bone strength, classical criteria for fluidity at the continuum level are often applied. Thus, the von Mises yield criterion is the most frequently used [11] criterion used for the trabecular bone of a vertebra, but a tissue-dependent orthotropic yield criterion Tsai–Wu [16] was proposed also.

In most cases, studies were performed by using static loads, while the number of investigations with dynamic loads are rather limited [13,17]. Also, the computational biomechanics can provide the ability to assess fracture risk and can serve as a useful diagnostic tool for analysing the state of osteoporotic bones. One of the most effective modern indicators used to verify vertebral fragility is the the trabecular bone score (TBS). This method involves the measurement of gray level texture, which shows the average state of trabecular bone microarchitecture [18], closely related to the ratio volume–total bone volume (BV/TV) [19]. Although TBS is a rough estimate of the strength properties of bone tissue, it can also be provided to study bone fragility and predict bone fractures.

There are several developed methods which employ bone mineral density (BMD) to predict the long-term fracture risk. BMD is usually assessed by dual-energy X-ray absorptiometry (DXA), and some studies have shown that a decrease of BMD is bonded with a higher risk for future fracture [20–22]. In addition, modern techniques of medical diagnostics also dispose computer-based algorithms, such as FRAX [23], which calculate fracture probability from clinical risk factors, such as age and body mass index (BMI), and dichotomized risk factors comprising prior fragility fracture, parental history of hip fracture, currently smoking tobacco, long-term oral glucocorticoid use, rheumatoid arthritis, other causes of secondary osteoporosis, and alcohol consumption [24]. In contrast, these methods do not verify the complicated relationship between the important parameters such as BV/TV, the thickness of the cortical shell, and the external load. The finite element method (FEM)-based continuum models can supply the additional patient-specific data to define the risk of fracture by additionally applying the reliability theory.

One of the aims of present work is numerical investigation of the osteoporotic affect on strength properties of the proposed vertebrae model including the time-dependent load. The evaluation results might be useful for the medical diagnostics of the osteoporosis and for verifying the strength properties of the bone tissue of the patient. The main aim of this study is proposing the structural mechanics-based method of calculating fracture risk based on statistically processed results obtained by the numerical investigation of strength properties of the lumbar vertebral L1 body with various grades of osteoporotic degradation.

2. Methods and Materials

2.1. Problem Formulation

In order to model bone tissue as elastoplastic three-dimensional solid, the Ramberg–Osgood stress–strain equation was applied [25]:

$$\varepsilon = \frac{\sigma}{E_0 \dot{\varepsilon}^d} + \alpha \left(\frac{\sigma}{K} \right)^n \dot{\varepsilon}^b \quad (1)$$

where E_0 is the modulus of elasticity, K is the strain hardening coefficient, n is the strain-hardening exponent, and α is the yield offset, where $\dot{\varepsilon} = 1$ equality was used. According to research [25], the loading rate of cortical bone influences the behaviour of the stress–strain relation. Therefore, our model of risk evaluation should be treated as the upper risk limit for load rates, which do not exceed rates of $\dot{\varepsilon} \leq 1$.

In this work, the von Mises–Hencky criterion is chosen to predict the fracture of the model. It is defined in Equation (2): There, σ_1 , σ_2 and σ_3 are the maximum, intermediate and minimum principal stresses. σ_Y presents yield stress (40 MPa) [26].

$$\sqrt{\frac{(\sigma_1 - \sigma_2)^2 + (\sigma_2 - \sigma_3)^2 + (\sigma_3 - \sigma_1)^2}{2}} = \sigma_Y \tag{2}$$

In the case of nonlinear analysis, the equilibrium equations are presented at time step $t + \Delta t$:

$$\mathbf{M}^{t+\Delta t} \mathbf{U}''^{(i)} + \mathbf{C}^{t+\Delta t} \mathbf{U}'^{(i)} + {}^{t+\Delta t} \mathbf{K}^{(i)} \Delta \mathbf{U}^{(i)} = {}^{t+\Delta t} \mathbf{R} - {}^{t+\Delta t} \mathbf{F}^{(i-1)} \tag{3}$$

where \mathbf{M} is the mass matrix, \mathbf{C} is the damping matrix, ${}^{t+\Delta t} \mathbf{K}^{(i)}$ is the stiffness matrix, ${}^{t+\Delta t} \mathbf{R}$ is the vector of nodal loads, ${}^{t+\Delta t} \mathbf{F}^{(i-1)}$ is the vector of nodal forces in case of iteration iteration $(i - 1)$, ${}^{t+\Delta t} \Delta \mathbf{U}^{(i)}$ is the vector nodal displacements while the iteration is (i) , ${}^{t+\Delta t} \mathbf{U}'^{(i)}$ is the vector of total velocities while the iteration is (i) , and $\mathbf{M}^{t+\Delta t} \mathbf{U}''^{(i)}$ is the vector of total accelerations while the iterations is (i) . Using an implicit time integration Newmark–Beta scheme and employing Newton’s iterative method, the presented equations are cast in the following form:

$${}^{t+\Delta t} \mathbf{K}^{(i)} \Delta \mathbf{U}^{(i)} = {}^{t+\Delta t} \mathbf{R} \tag{4}$$

where ${}^{t+\Delta t} \mathbf{R}^{(i)}$ is the effective load vector and ${}^{t+\Delta t} \mathbf{K}^{(i)}$ is the effective stiffness matrix. The three-dimensional static and dynamic analyses were performed using Abaqus (c) [27] software.

2.2. Structure and Geometrical Properties of the Model

The inhomogeneous model is made of two basic parts members: the lumbar body with posterior elements and intervertebral disks. The lumbar geometrical form is extracted by processing DICOM format data and then converted into a numerical model using Slice3D software. It is presented in Figure 1a.

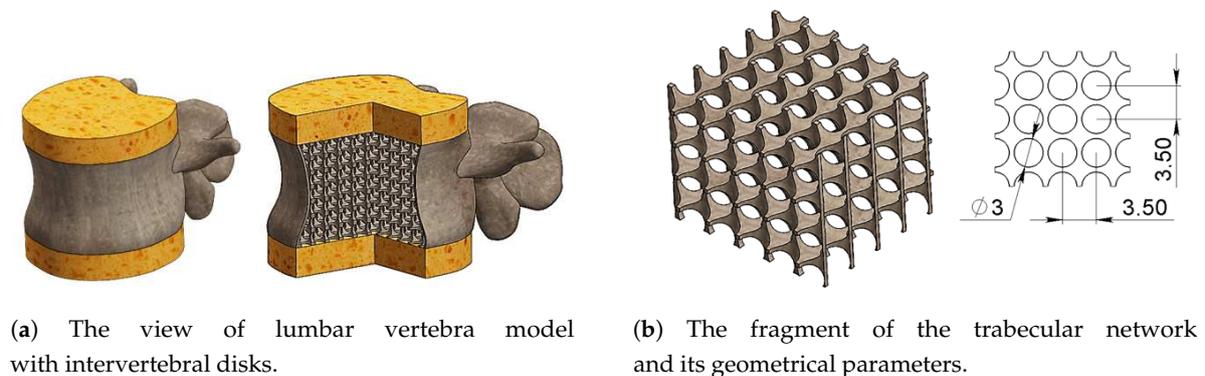


Figure 1. Geometrical models of vertebra.

The trabecular network is formed by cylindrical cuts which are extruded from vertical and horizontal planes. This method allowed the creation of the characteristic structure made of rod-like and plate-like trabeculas, and a fragment of the trabecular lattice is shown in Figure 1b.

2.3. Mechanical Properties

The nonlinear stress–strain diagram of bone is presented in Figure 2. On the other hand, the intervertebral disks were considered as isotropic and perfectly elastic. The material constants of the listed components were taken from data reported by Reference [10] and presented in table Table 1.

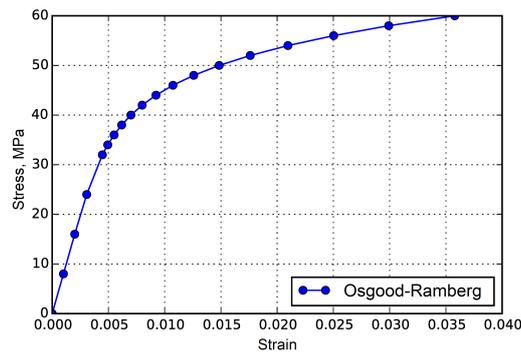


Figure 2. Osgood–Ramberg stress–strain relation for cortical bone.

Table 1. Elasticity constants of model components.

Property	Lumbar Body	Intervertebra Disc
Young modulus, MPa	8000	10
Ultimate strength, MPa	60	-
Yield strength, MPa	40	-
Poisson’s ratio	0.30	0.40

2.4. The Parameters of Osteoporosis Impact

The effect of osteoporotic degradation is implemented by thickening the cortical bone and by varying the TBS, which depending on BV/TV. This research proposes the investigation of three parametric models with various cortical thicknesses of 0.2, 0.4, and 0.5 mm first offered in References [12,14]. Also, the relation of the BV/TV as well as the TBS value is typical for a high level of osteoporosis. The data taken from References [18,19] is presented in Table 2.

Table 2. The parameters of a healthy and osteoporotic model.

Model	Thickness, mm	BV/TV	TBS
Healthy	0.5	0.35	1.45
Ostseopenea	0.4	0.20	1.33
Osteoporosis	0.2	0.10	1.20

To make a profound investigation, the parameters in Table 2 were combined in different cases of calculations that were performed for both static and dynamic analyses (Table 3).

Table 3. The parameters of different lumbar body models.

Cortical Shell Thickness, mm	BV/TV	TBS
0.20	0.10	1.28
0.40	0.10	1.28
0.50	0.10	1.28
0.20	0.20	1.33
0.40	0.20	1.33
0.50	0.20	1.33
0.20	0.35	1.45
0.40	0.35	1.45
0.50	0.35	1.45

2.5. Boundary Conditions and Mesh

The model was tested by the external loads in the range of 0.15–0.75 MPa which arise in the results of daily motions [14]. The conditions of loading are reflected in Figure 3. Also, in the case of the nonlinear problem, the load direction is bonded with the displacement values and, during the test, follows the deformed shape of the model.

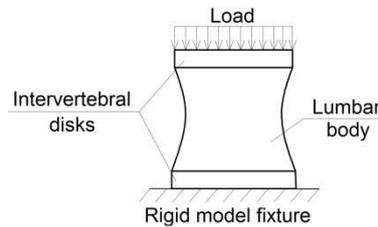


Figure 3. Load conditions of the compression test.

The time curve of the load is presented in Figure 4 and simulates the bearing of dynamic load. The time curve character comes from experimental data reported by Reference [28].

The model is meshed with tetrahedral finite elements and characterized by 306,435, 277,896, and 256,438 mesh elements for BV/TV ratios of 0.35, 0.20, and 0.1 or TBS values of 1.45, 1.33, and 1.2, respectively.

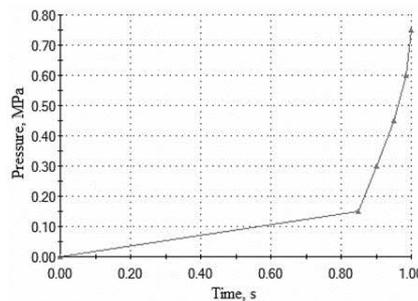


Figure 4. Time curve of dynamic load.

The contact between vertebra and intervertebral disks is treated as no penetration-bonded contact.

2.6. Risk Evaluation

The following three random variables were used to evaluate risk: external load P , the cortical shell thickness of lumbar vertebra Δ , and bone volume to total volume ratio $\beta_{BV/TV}$ which was obtained from TBS interdependence. In contrast, TBS was obtained from the 2-D vertebra CT scan. These variables were used for solving the FEM for the proposed mechanical model strength of the lumbar vertebra. Finally, the obtained set of points values for different combinations of independent variables were used for least square fitting by cubic polynomial as follows

$$\sigma_s(\beta, \delta, p) = \sum_{i,j,k=0}^{i+j+k \leq 3} a_{ijk} \beta^i \delta^j p^k \tag{5}$$

where p is the external load, δ is the cortical shell thickness, and β is the BV/TV ratio.

Furthermore, the approximation was obtained by using logarithmic values of FEM stresses $\log(\sigma_{vonMises})$ by searching minima of the following expression:

$$\min \left| \sum_{i,j,k} \log(\sigma_{vonMises}(i, j, k)) - \sigma_s(\beta_i, \delta_j, p_k) \right| \tag{6}$$

where indexes i, j, k denote values of discrete independent variables used in FEM computations.

Fracture risk F_R of lumbar vertebra can be expressed in term of reliability:

$$R = P(Z \leq 0) = 1 - F_R(Z \leq 0), \tag{7}$$

where $Z = X_\sigma - Y_{\sigma_Y}$, X_σ is max stress, and Y_{σ_Y} is yield strength. The general expression for the fracture risk of a stress–strength system has the following cumulative density function (CDF) [29]:

$$F_R = 1 - P(Z \leq 0) = 1 - \int_0^\infty \int_0^\infty f_{X_\sigma}(\sigma + z) f_{Y_{\sigma_Y}}(\sigma) d\sigma dz \tag{8}$$

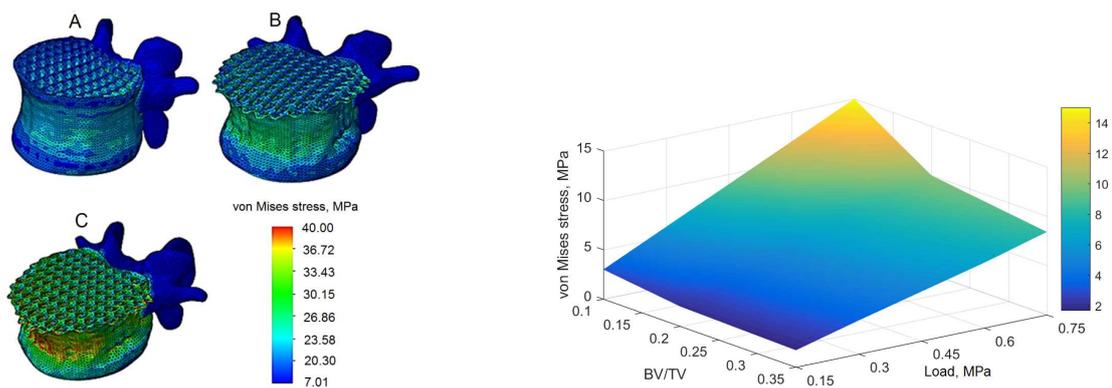
where f_{X_σ} and $f_{Y_{\sigma_Y}}$ are probability density functions (PDFs) of the random values X_σ and Y_{σ_Y} of the maximum stress caused by the external load p and the strength, respectively.

3. Numerical Results and Discussion

Important parameters that impact the state of lumbar body, such as the thickness of the cortical shell and the BV/TV relation, were combined between themselves and investigated due to various values of external load. Also, the numerical results of both static and dynamic analysis and the discussion of the results are presented in the following section.

3.1. Static Analysis

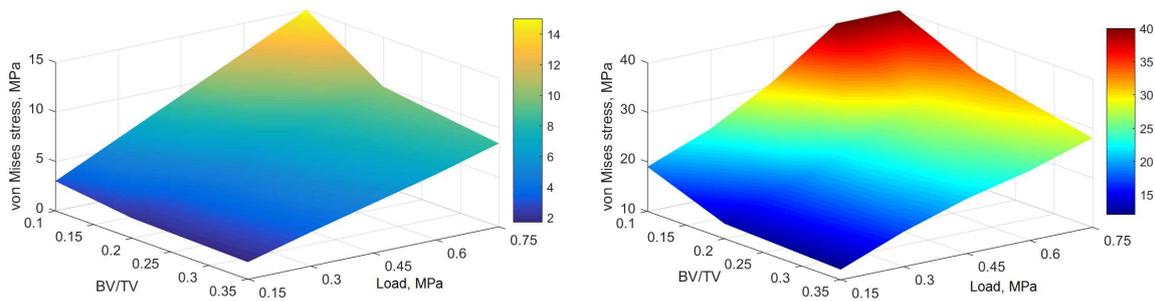
Contours of the deformed shape of the model were generated, and stress distribution on the model is shown in Figure 5. The figure shows that the highest von Mises stresses occurred in the middle of the cortical shell on the front side of the model and that the deformations are scaled by a factor of 20 for visualization purposes. Figure 5b shows a relation between the von Mises stress versus the BV/TV ratio and the external load. The results show that maximal calculated stress reaches 38% of yield stress due to a 0.75 MPa load on the model with the lowest BV/TV ratio (0.1). Figure 5c,d shows the von Mises stress for models with thicknesses of cortical shell of 0.4 and 0.2 mm, respectively. In the case of a cortical shell with a thickness of 0.4 mm, maximal stress is about 17 MPa due to the minimal BV/TV ratio and the maximum external load (0.75 MPa), reaching 43% of the yield stress.



(a) Von Mises stress on the vertebra model: (A) when BV/TV is 35% and cortical shell thickness is 0.5 mm; (B) when BV/TV is 20% and the cortical shell thickness is 0.4 mm; and (C) as BV/TV is 10% and the cortical shell thickness is 0.2 mm. The cortical shell thickness is 0.5 mm

(b) A relation between the von Mises stress versus the BV/TV ratio and the external load. The thickness of cortical shell is 0.4 mm

Figure 5. Cont.



(c) Maximum von Mises stress versus the BV/TV ratio and the external load. The thickness of cortical shell is 0.5 mm

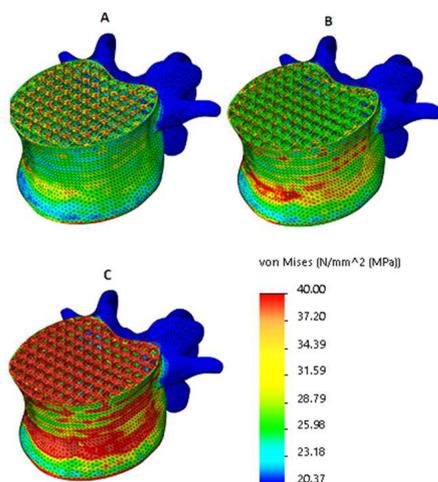
(d) Maximum von Mises stress versus the BV/TV ratio and the external load. The thickness of cortical shell is 0.2 mm

Figure 5. Results of the static analysis.

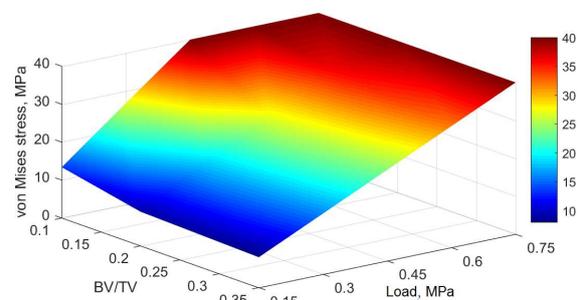
Figure 5d shows that the strength load is 0.60 MPa (the yield stress is exceeded), while the BV/TV ratio is 0.1. The obtained results agree well with the research of Kim et al., [14], with a difference about 5–10% for the model with a low apparent density and a cortical shell thickness of 0.2 mm, while McDonald [12] reports this value up to 0.99 MPa.

3.2. Dynamic Analysis

The stress distribution in the lumbar vertebra during dynamical analysis was computed, and the analysis shows that the stress concentrators have greater differences than seen in the static analysis. The highest value of stress MPa appeared in the middle of the cortical shell and agrees well with clinical observations [1–4,30]. Also, the distribution of the von Mises stress on the cortical shell of the model is shown in Figure 6a. Figure 6b–d presents a relation of the von Mises stress on the model versus the BV/TV ratio and the external load. Figure 6d shows that yield stress is exceeded due to the 0.45 MPa load while the BV/TV ratio is less than 0.1. The relations between the von Mises stress on the cortical shell of the model with thicknesses of 0.4 and 0.2 mm versus the BV/TV ratios and the external load are presented in Figure 6c,d, respectively.

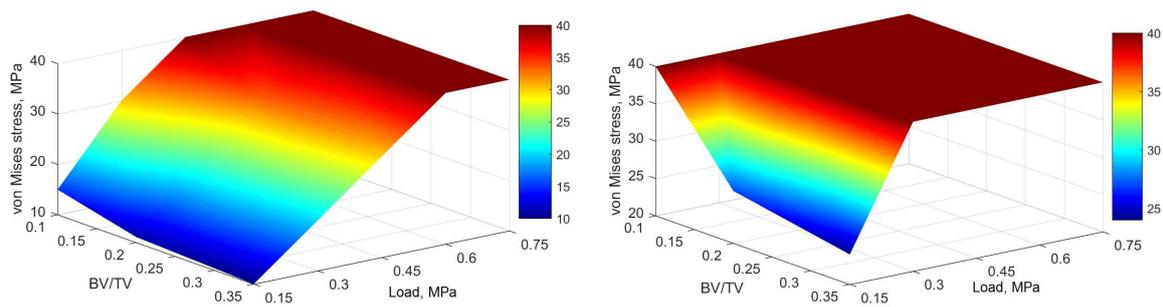


(a) Von Mises stress on the vertebra model: (A) when BV/TV is 35% and cortical shell thickness is 0.5 mm; (B) when BV/TV is 20% and the cortical shell thickness is 0.4 mm; and (C) as BV/TV is 10% and the cortical shell thickness is 0.2 mm



(b) Maximum von Mises stress versus the BV/TV ratio and the external load. The cortical shell thickness is 0.5 mm

Figure 6. Cont.



(c) Maximum von Mises stress versus the BV/TV ratio and the external load. The cortical shell thickness is 0.4 mm

(d) Maximum von Mises stress versus the BV/TV ratio and the external load. The cortical shell thickness is 0.2 mm

Figure 6. Results of the dynamic analysis.

As shown in Figure 6c, the excess yield stress is expected due to the 0.45 MPa load, while the BV/TV ratio is minimal. In the case of a high BV/TV ratio, the excess yield stress was identified due to the load of 0.6 MPa in terms of the average thickness of the cortical shell (0.4 mm). Figure 6d shows that yield stress is exceeded almost in every sample of bone tissue if the load is higher than 0.3 MPa. In addition, if the load grows increases 0.3 MPa, the yield stress is reached even in the case of a sufficient BV/TV ratio (0.35). This effect can be explained by the low thickness of the cortical shell (0.2 mm).

3.3. Estimation of the Fracture Risk

The risk of fracture of the lumbar vertebra can be estimated by the Monte Carlo method as a probability $Pr_f = Pr(Z \leq 0)$, where $Z = Y_{p,R} - X_{p,E}$ and $X_{p,E}$ and $Y_{p,R}$ are external pressure p imposed on the vertebra and the load-bearing capacity of the vertebra random variables respectively. It is assumed that the load-bearing capacity of the vertebra p_R equals the minimum pressure p at which the maximum von Mises stresses equal the yield stress σ_Y , that is, for the given δ and β , the load-bearing capacity r.v. $p_R = \min\{p : g_{FEM}(\delta, \beta, p) \geq \sigma_Y\}$. It is assumed that $X_{p,E}$ and $Y_{p,R}$ are independent random variables. To evaluate the fracture risk Pr_f , the random numbers of the acting pressure random variable $X_{p,E}$ can be generated directly, while the generation of the random numbers of the load-bearing capacity $Y_{p,R}$ is more complicated since it is known that the cortical shell thickness and BV/TV are dependent random variables. In addition, the load-bearing capacity function of the lumbar vertebra $p_R = h(\delta, \beta, \sigma_Y, \dots)$ dependent of all variables must be known. Then, the r.v. of the load-bearing capacity $Y_{p,R}$ can be obtained as a transformation $Y_{p,R} = h(X_\delta, X_\beta, X_{\sigma_Y}, \dots)$, where $X_\delta, X_\beta, X_{\sigma_Y}, \dots$ are random variables affecting the load-bearing capacity of the lumbar vertebra: for example, the cortical shell thickness, BV/TV ratio, the yield stress, geometry, and so on. It is evident that it is not possible to generate the sample of the realizations of the load-bearing capacity r.v. $Y_{p,R}$ directly by calculating each random number by FEM. Therefore, in the present article, another approach is adopted: to approximate the maximum von Mises stresses of the vertebra by the polynomial \hat{g} and then, by solving this polynomial as an equation $\sigma_Y = \hat{g}(\delta, \beta, p)$, to obtain the pressure p that corresponds to the given variables δ, β , and the maximum von Mises stress σ_Y .

3.3.1. Approximation of the Maximum von Mises Stress

The approximation of the calculated maximum von Mises stresses of the lumbar vertebra was conducted by using the program “R studio” [31] and the package “nlme” [32] over the set $\mathbb{D} \times \mathbb{B} \times \mathbb{P}$, where \mathbb{D}, \mathbb{B} , and \mathbb{P} are sets of values of the cortical shell thickness δ , BV/TV ratio β , and the external load p , respectively: $\delta \in \mathbb{D} = \{0.5, 0.4, 0.2\}$ mm, $\beta \in \mathbb{B} = \{0.1, 0.2, 0.35\}$, and $p \in \mathbb{P} = \{0.15, 0.30, 0.45, 0.60, 0.75\}$ MPa. The set of 45 points of the maximum von Mises stresses

$\mathbb{D} \times \mathbb{B} \times \mathbb{P}$ of the static analysis obtained by the finite element method that were used to obtain the approximating polynomial is given in Table 4.

Table 4. The maximum von Mises stresses of the static analysis calculated by finite element method (FEM) at different δ , β , and p

Analysis	Thickness, mm, δ	BV/TV, β	External Load, MPa, p				
			0.15	0.30	0.45	0.60	0.75
Static	0.5	0.10	2.9	6.1	9.3	12.3	15.2
		0.20	2.0	4.1	6.1	8.2	8.5
		0.35	1.6	3.4	5.0	6.7	8.5
	0.4	0.10	3.4	6.7	10.2	14.1	17.0
		0.20	2.1	4.2	6.3	8.4	10.5
		0.35	1.8	3.5	5.4	7.3	9.1
	0.2	0.10	18.9	24.1	31.3	40.8	49.4
		0.20	13.3	18.4	23.2	28.1	33.5
		0.35	12.1	17.4	21.0	24.5	28.0

The approximating polynomial of the maximum von Mises stresses of the static analysis is as follows:

$$\sigma \approx \hat{g}(p, \beta, \delta) = a_0 + a_1p + a_2\beta + a_3\delta + a_{11}p^2 + a_{33}\delta^2 + a_{12}p\beta + a_{13}p\delta + a_{123}p\beta\delta + a_{112}p^2\beta + a_{113}p^2\delta + a_{221}\beta^2p + a_{223}\beta^2\delta + a_{331}\delta^2p + a_{332}\delta^2\beta + a_{12233}p\beta^2\delta^2 + a_{11233}p^2\beta\delta^2, \delta \in [0.2, 0.5] \text{ mm}, \beta \in [0.1, 0.35], p \in [0.15, 0.75] \text{ MPa} \quad (9)$$

where the coefficients are as follows: $a_0 = 40.095, a_1 = 109.545, a_2 = -33.064, a_3 = -153.799, a_{11} = 56.976, a_{33} = 152.808, a_{12} = -410.969, a_{13} = -348.377, a_{123} = 704.586, a_{112} = -153.953, a_{113} = -117.969, a_{221} = 737.456, a_{223} = 97.428, a_{331} = 361.940, a_{332} = 56.473, a_{12233} = -2913.652, a_{11233} = 702.954.$

It should be noted that only the coefficients a_i for which the statistical hypothesis $H_0 : a_i = 0$ is rejected and the alternative $H_1 : a_i \neq 0$ is accepted are present in the given approximating polynomial Equation (9). The hypotheses were tested using the Student t -test. The maximum and minimum residuals are as follows: $\max\{g_{FEM}(p_i, \beta_i, \delta_i) - \hat{g}(p_i, \beta_i, \delta_i)\} = 0.883$ MPa and $\min\{g_{FEM}(p_i, \beta_i, \delta_i) - \hat{g}(p_i, \beta_i, \delta_i)\} = -0.838$ MPa, where $(p_i, \beta_i, \delta_i) \in \mathbb{D} \times \mathbb{B} \times \mathbb{P}$ and $g_{FEM}(p_i, \beta_i, \delta_i)$ stand for the stress values calculated by the FEM. The relative residuals $(\hat{g}(p_i, \beta_i, \delta_i) - g_{FEM}(p_i, \beta_i, \delta_i)) / g_{FEM}(p_i, \beta_i, \delta_i)$ versus the stresses calculated by FEM are depicted in Figure 7.

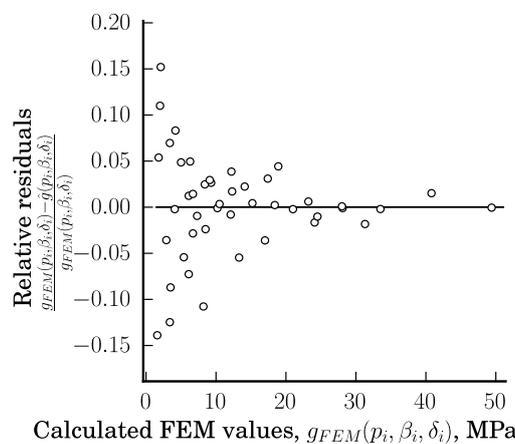


Figure 7. Relative residuals versus calculated values of the maximum von Mises stress.

The obtained polynomial \hat{g} , given in Equation (9), can be treated as a quadratic equation $\sigma_Y = \hat{g}(p, \beta, \delta)$ with respect to the pressure p when the other two variables δ and β are treated as known quantities or, in other words, as the parameters. Then, the solution of the obtained polynomial \hat{g} , given in Equation (9), gives two roots: One of them is the load-bearing capacity p_R corresponding to the given δ, β , and the yield stress σ_Y :

$$p_R \approx \hat{h}(\beta, \delta) = \frac{-b_1(\beta, \delta) + \sqrt{b_1(\beta, \delta)^2 - 4b_2(\beta, \delta)(b_0(\beta, \delta) - \sigma_Y)}}{2b_2\beta, \delta} \tag{10}$$

where $b_0(\beta, \delta) = a_{332}\beta\delta^2 + a_{333}\delta^2 + a_{223}\beta^2x_3 + a_3\delta + a_2\beta + a_0$, $b_1(\beta, \delta) = a_{12233}\beta^2\delta^2 + a_{331}\delta^2 + a_{123}\beta\delta + a_{13}\delta + a_{221}\beta^2 + a_{12}\beta + a_1$ and $b_2(\beta, \delta) = a_{11233}\beta\delta^2 + a_{113}\delta + a_{112}\beta + a_{11}$.

Then, the load-bearing capacity r.v. $Y_{p,R}$ can be modelled as the transformation $Y_{p,R} = \hat{h}(X_\beta, X_\delta)$ of r.vs. X_β and X_δ by assuming that the yield stress σ_Y is the deterministic predefined quantity.

3.3.2. Stochastic Models and for the Risk Modelling

It is well-known that osteoporosis of the trabecular bone also entails decreasing of the thickness of the cortical shell of the lumbar vertebra δ . Therefore, BV/TV and the cortical shell thickness are dependent random variables. In addition, decreasing BV/TV also entails decrease of the cortical shell thickness δ [33,34]. Therefore, the correlation coefficient of these r.vs. is positive. For example in Reference [33], it is obtained that the statistically significant coefficient of correlation between the cortical shell thickness and BV/TV is 0.5, when the significance level is 0.05. According to Reference [34], the coefficient of correlation between BV/TV and the cortical shell thickness is 0.29; however, it is statistically insignificant. To generate the load-bearing capacity $X_{p,R}$ random numbers, it is assumed that the BV/TV ratio and the cortical shell thickness form a bivariate correlated truncated Gaussian random variable $(X_\beta, X_\delta)^T \in [M(G_\beta) - 3\sqrt{D(G_\beta)}, M(G_\beta) + 3\sqrt{D(G_\beta)}] \times [M(G_\delta) - 3\sqrt{D(G_\delta)}, M(G_\delta) + 3\sqrt{D(G_\delta)}]$; where \times is the Cartesian product of two sets; $M(G_\beta)$, $D(G_\beta)$, $M(G_\delta)$, and $D(G_\delta)$ are the means and variances of the correlated untruncated normal random variables G_β and G_δ , respectively, that also form a bivariate Gaussian random variable $(G_\beta, G_\delta)^T \sim N(\vec{\mu}, \Sigma)$ of which the mean vector is $\vec{\mu} = (M(G_\beta), M(G_\delta))^T$ and the covariance matrix is

$$\Sigma = \begin{bmatrix} D(G_\beta) & Cor(G_\beta, G_\delta)\sqrt{D(G_\beta)D(G_\delta)} \\ Cor(G_\delta, G_\beta)\sqrt{D(G_\delta)D(G_\beta)} & D(G_\delta) \end{bmatrix} \tag{11}$$

where $Cor(G_\beta, G_\delta)$ and $Cor(G_\delta, G_\beta)$ are the correlation coefficients between r.vs. G_β and G_δ , $Cor(G_\beta, G_\delta) = Cor(G_\delta, G_\beta)$.

R.v. $Y_{p,R}$ can be estimated by the empirical cumulative distribution function (ECDF) and by the empirical quantile function (EQF). Usually, ECDF is defined as follows: $Pr(Y_{p,R} \leq y) \approx \hat{F}_{Y_{p,R}}(y) = \frac{1}{n} \sum_{i=1}^n I(y_i \leq y)$, where I is the indicator function, n is the sample size, y_i is the realization i of r.v. $Y_{p,R}$; while EQF of r.v. $Y_{p,R}$ is usually defined as $Q(p) = inf\{y : F_{Y_{p,R}}(y) \geq p\}$, where p is the required probability.

For the evaluation of the fracture risk as the probability $Pr_f = Pr(Y_{p,R} - X_{p,E} \leq 0)$, the external pressure independent r.v. $X_{p,E}$ can be distributed according to various laws. In the present article, it is assumed that the external load r.v. $X_{p,E}$ is also truncated normal r.v. attaining values from the interval $X_{p,E} \in [M(G_{p,E}) - 3\sqrt{D(G_{p,E})}, M(G_{p,E}) + 3\sqrt{D(G_{p,E})}]$, where $M(G_{p,E})$ and $D(G_{p,E})$ are the mean and the variances of the untruncated external pressure normal r.v. $G_{p,E}$.

The fracture risk of the lumbar vertebra can be expressed as $Pr_f = Pr(Y_{p,R} - X_{p,E} \leq 0) = Pr(Z \leq 0) \approx \hat{F}_Z(0)$, where \hat{F}_Z is the ECDF of r.v. $Z = Y_{p,R} - X_{p,E}$: $\hat{F}_Z(z) = \frac{1}{n} \sum_{i=1}^n I(y_{p,R,i} - x_{p,E,i} \leq z)$, where n is the sample size and $y_{p,R,i}$ and $x_{p,E,i}$ are realization i

of r.v.s. $Y_{p,R}$ and $X_{p,E}$, respectively. It should be noted that the fracture risk can be also evaluated according to the following known integral [29]:

$$Pr_f = P(Z \leq 0) = \int_{-\infty}^0 \int_{-\infty}^{+\infty} f_{Y,p,R}(p+z)f_{X,p,E}(p)dpdz \tag{12}$$

where $f_{Y,p,R}$ and $f_{X,p,E}$ are probability density functions (PDFs) of the load-bearing capacity r.v. $Y_{p,R}$ and the external pressure r.v. $X_{p,E}$, respectively.

For the following analysis of the reliability of the vertebra, the concrete parameters of the r.v. G_β , G_δ , and $G_{p,E}$ must be adopted. In the case of the Gaussian r.v.s., only the means, variances, and covariances are required. The coefficient of variation $Cvar(G_\beta)$ of BV/TV can vary in a wide range, for example, from about 18% [18,33] up to 37% [34] or even up to 42% for the normal bones [26]. There is less literature on the coefficient of the variation of the cortical shell thickness δ . We found that the coefficient of variation is also big, for example, 24.5% [33] or even 62.4%, according to Reference [34]. For the patient-specific cases, the variability of these parameters should be less. However, since the BV/TV ratio is determined indirectly, the variability of r.v. X_β cannot be very low.

As it was already mentioned, the r.v.s. X_β and X_δ are dependent [33,34]. In the case of the Gaussian bivariate r.v. (G_β, G_δ) , the correlation coefficient $Cor(G_\delta, G_\beta)$ is used in the covariance matrix given in Equation (11) and it is determined that $Cor(G_\delta, G_\beta)$ may equal 0.5 [33] or 0.29 [34]. The external load $X_{p,E}$ acting on the vertebra is a random variable depending mainly on the mechanical loads. For example, according to Reference [28], the ground reaction of a human may vary from 750 N in the rest state and increases up to 2000 N in the squat jump. The ground reaction of the rest state shows that the mass of the investigated human body in Reference [28] is about 75 kg. Then, in the squat jump, the weight of the body increases $2000/750 = 2.66(7)$ times. It should be mentioned that not all the weight of a body is imposed on the lumbar vertebra. On the basis of the data given in Reference [35], it is possible to conclude that a lumbar vertebra sustains about 60% of the total load. By taking into account the area of the loaded surface of the vertebra under investigation, $A = 1210 \text{ mm}^2$, we can conclude that, for the person of article [28], the external load varies within an interval $[p_{min}, p_{max}] = [0.372, 0.992]$ MPa.

3.3.3. Fracture Risk Modelling Data, Results, and Discussion

On the basis of the above given review, for the probabilistic estimations of the load-bearing capacity of the lumbar vertebra $Pr(Y_{p,R} \leq y_p) = \hat{F}_{Y,p,R}(y_p)$ and the failure risk $Pr_f = Pr(Z \leq 0) = F_Z(0)$, the following two cases were considered. For the first case (Case I), the following values are adopted: the mean of BV/TV $M(G_\beta) = 0.1$ and the mean of the cortical shell thickness $M(G_\delta) = 0.2 \text{ mm}$; the coefficients of variation $Cvar(G_\beta) = \sqrt{D(G_\beta)}/M(G_\beta) = 0.2$ and $Cvar(G_\delta) = 0.2$; the correlation coefficient $Cor(G_\beta, G_\delta) = 0.35$; the mean of the external load $G_{p,E}$ attains three values $M(G_{p,E}) \in \{p_{min}, 1.5p_{min}, 2.5p_{min}\}$, where $p_{min} = 0.372 \text{ Mpa}$; and the coefficient of variation for all external pressures is constant: $Cvar(G_{p,E}) = 0.20$. For the second case (Case II), the following values are adopted: means $M(G_\beta) = 0.25$ and $M(G_\delta) = 0.2 \text{ mm}$; the coefficients of variation $Cvar(G_\beta) = 0.2$ and $Cvar(G_\delta) = 0.2$; the correlation coefficient $Cor(G_\beta, G_\delta) = 0.35$; mean of the external load $G_{p,E}$ attains three values $M(G_{p,E}) \in \{p_{min}, 1.5p_{min}, 2.5p_{min}\}$, where $p_{min} = 0.372 \text{ Mpa}$; and the coefficient of variation for all external pressures is constant: $Cvar(G_{p,E}) = 0.2$. Thus, for both Cases I and II, the external load $G_{p,E}$ attains the same three different means and the coefficient of variation is constant, and for both cases, the coefficient of the correlation between the cortical shell thickness and BV/TV is the same, 0.35. It should be noted that for Cases I and II the mean of the external pressure $M(G_{p,E}) = 0.372 \text{ MPa}$ approximately corresponds to the external load of the lumbar vertebra of a person of mass 75 kg in rest, while $M(G_{p,E}) = 2.5p_{min} = 0.93 \text{ MPa}$ is close to the maximum load during the squat jump of the person. Totally, 10^6 random values of the r.v.s. G_β and G_δ were generated. However, due to the truncation the

samples of the realizations of the r.v.s. X_β , X_δ , $X_{p,E}$, and $Y_{p,R}$ were smaller than 10^6 but bigger than 9×10^5 . All analyses were performed by the program “R studio” [31]. The bivariate Gaussian random numbers were generated using the package “MASS” [36].

The realizations of the histograms of the r.v.s. X_β , X_δ and $Y_{p,R}$ are shown in Figure 8. The realizations of the histograms of the truncated BV/TV ratio r.v. X_β are depicted in Figure 8a,d and the realizations of the truncated cortical shell thickness r.v. X_δ are depicted in Figure 8b,e while realizations of the histograms of the load-bearing capacity r.v. $Y_{p,R}$ as well as the 0.05 and 0.95 empirical quantiles of r.v. $Y_{p,R}$ are depicted in Figure 8c,f. It should be noticed that the histograms depicted in Figure 8a,b,d,e show entire ranges of the random numbers of the r.v.s. X_β and X_δ , while Figure 8c,d show truncated ones from the right of the intervals of r.v.s. $Y_{p,R}$.

As we can see from Figure 8c,f, the probability density function of the load-bearing capacity r.v. $Y_{p,R}$ can have a positive skew. However, there are such cases of $M(X_\beta)$, $M(X_\delta)$, $D(X_\beta)$, and $D(X_\delta)$ in which the histograms of r.v. $Y_{p,R}$ have a negative skew. It was also observed that at the small values of variances $D(X_\beta)$ and $D(X_\delta)$, the histograms of the r.v. $Y_{p,R}$ are symmetrical and the histograms are very similar to the normal distribution probability density function with mean $M(Y_{p,R})$ and variance $D(Y_{p,R})$. However, the given Figure 8c,f shows a counterexample, and in general, we cannot state that r.v. $Y_{p,R}$ is normal provided that r.v.s. X_β and X_δ are symmetrically truncated dependent normal random variables with nonnegative correlation coefficient. As we can see from Figure 8c,f, the 0.05 quantiles of the load-bearing capacity of the vertebra r.v. $Y_{p,R}$ increase from $p_{0.050} = 0.355$ MPa for Case I up to $p_{0.05} = 0.638$ MPa for Case II. Analogously, the 0.95 quantiles increase from $p_{0.95} = 0.913$ MPa up to $p_{0.95} = 2.364$ MPa. The quantiles of the load-bearing capacity r.v. $Y_{p,R}$ can serve as extra information on the fracture risk of the lumbar vertebra. It should be noted that the correlation between r.v.s. X_β and X_δ in general is important for r.v. $Y_{p,E}$. For example, for the considered Case I, the increase of the correlation coefficient $Cor(G_\beta, G_{delta})$ from 0.35 up to 0.5 decreases the small quantiles of the load-bearing capacity r.v. $Y_{p,R}$ from $p_{0.05} = 0.355$ up to $p_{0.05} = 0.347$ and increases the big quantiles from $p_{0.05} = 0.913$ up to $p_{0.95} = 0.929$. For Case II, the increase the correlation coefficient from $Cvar(G_\beta, G_\delta) = 0.35$ up to $Cvar(G_\beta, G_\delta) = 0.5$ also entails a decrease of small value quantiles of r.v. $Y_{p,R}$ from $p_{0.05} = 0.638$ up to $p_{0.05} = 0.620$, but the increase in the correlation coefficient also increases the big quantile values from $p_{0.95} = 2.364$ up to $p_{0.95} = 2.423$.

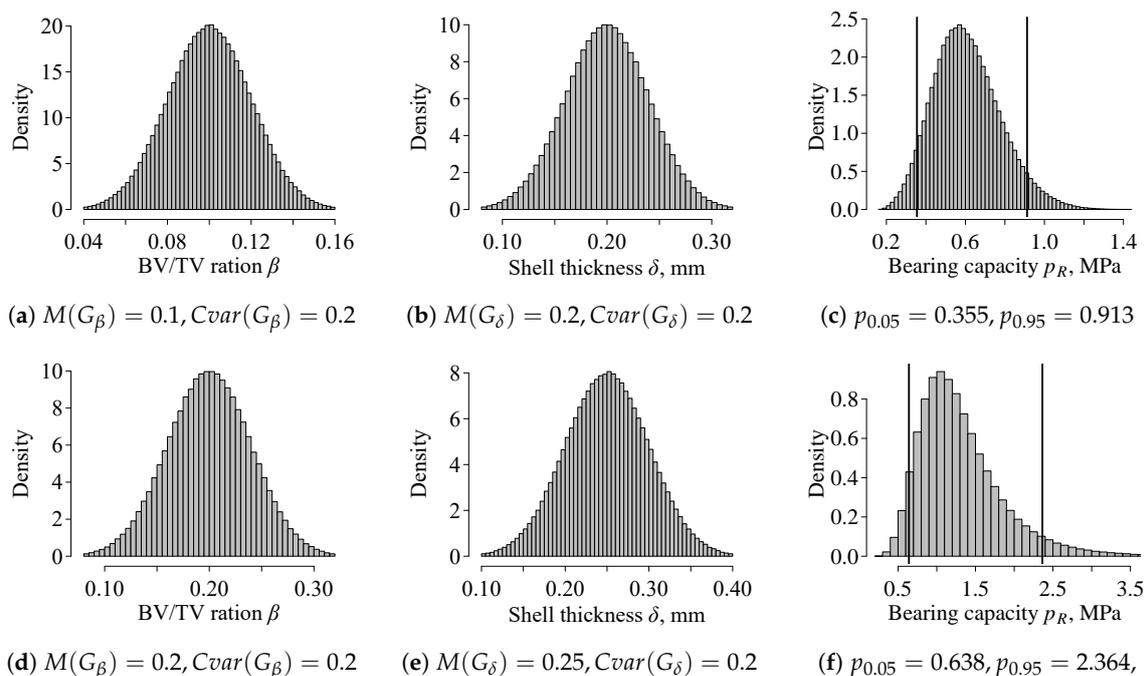


Figure 8. Realizations of the probability density functions of the random variables X_β (a,d); r.v. X_δ (b,e); and r.v. $Y_{p,R}$ (c,f).

The fracture risks of static vertebra loads of Case I and Case II as well histograms of r.v.s. $Y_{p,R}$, $X_{p,E}$, and $Z = Y_{p,R} - X_{p,E}$ can be seen in Figure 9. It was obtained that with increasing the cortical shell thickness from $\delta = 0.2$ (Case I) to $\delta = 0.25$ (Case II) and the BV/TV ratio from $\beta = 0.1$ (Case I) to $\beta = 0.2$ (Case II), the failure risk Pr_f decreases from 0.0904 to $2.752 \cdot 10^{-3}$ for $M(G_{p,E}) = p_{min} = 0.372$ MPa; from 0.417 to 0.0337 for $M(G_{p,E}) = 1.5p_{min} = 0.558$ MPa; and from 0.897 to 0.260 for $M(G_{p,E}) = 2.5p_{min} = 0.93$ MPa. The obtained results suggest that, for the considered supposed person, the squat jump inevitably would lead to the fracture of a lumbar vertebra, while a relatively small increasing BV/TV and the cortical shell thickness δ significantly decrease the fracture risk Pr_f .

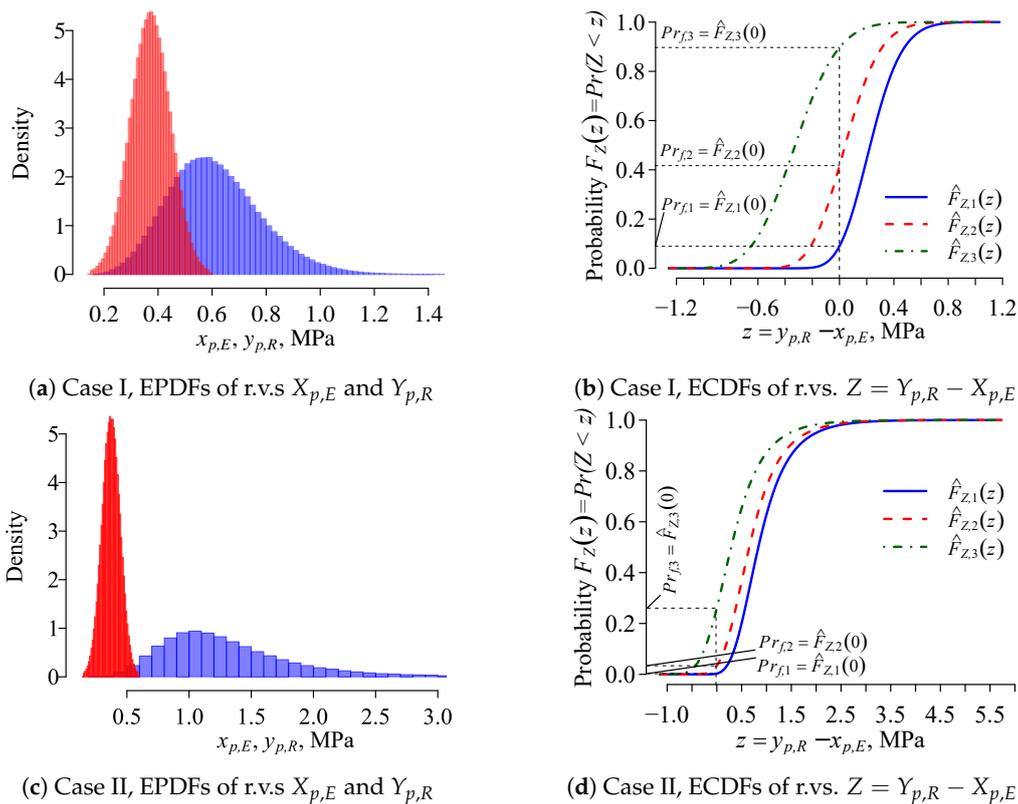


Figure 9. Realizations of the probability density functions of the random variables $X_{p,E}$ and $Y_{p,R}$ (a) for Case I and (c) for Case II; and empirical cumulative distribution functions $\hat{F}_{Z,i}(z)$ of r.v. $Z = Y_{p,R} - X_{p,E}$ and failure risks $Pr_{f,i}$, $i \in \{1, 2, 3\}$ for Case I (b) and for Case II (d): $\hat{F}_{Z,1}(z)$ and $pr_{f,1}$ when $M(G_{p,E}) = p_{min}$; $\hat{F}_{Z,2}(z)$ and $pr_{f,2}$ when $M(G_{p,E}) = 1.5p_{min}$; $\hat{F}_{Z,3}(z)$ and $pr_{f,3}$ when $M(G_{p,E}) = 2.5p_{min}$, where correlation coefficient $Cvar(G_{p,E}) = 0.2$ for all $Pr_{f,i}$, $\hat{F}_{Z,i}(z)$, $i \in \{1, 2, 3\}$.

The correlation between r.v.s. X_β and X_δ is also important for the fracture risk Pr_f . For example, for the considered Case I, the increase of the correlation coefficient $Cor(G_\beta, G_\delta)$ from 0.35 up to 0.5 increases the fracture risk from $Pr_f = 0.0904$ up to $Pr_f = 0.0963$ and from $Pr_f = 0.417$ up to $Pr_f = 0.418$ at small values of the loading, i.e., as $M(G_{p,E}) \in \{p_{min}, 1.5p_{min}\} = \{0.372, 0.558\}$ MPa. However, at big load values, the increasing correlation coefficient decreases the fracture risk from $Pr_f = 0.897$ up to $Pr_f = 0.892$ as $M(G_{p,E}) = 2.5p_{min} = 0.93$ MPa. For Case II, the increase in the correlation coefficient from $Cvar(G_\beta, G_\delta = 0.35)$ up to $Cvar(G_\beta, G_\delta) = 0.5$ increases the fracture risk from $Pr_f = 0.002752$ up to $Pr_f = 0.00387$ as $M(G_{p,E}) = p_{min} = 0.372$ MPa; from $Pr_f = 0.0337$ up to $Pr_f = 0.0392$ as $M(G_{p,E}) = 1.5p_{min} = 0.558$ MPa; and from $Pr_f = 0.260$ up to $Pr_f = 0.270$ as $M(G_{p,E}) = 2.5p_{min} = 0.93$ MPa.

It should be noted that the probabilistic evaluation of the fracture risk demands information on the probability distribution functions of the many variables that affect the failure of a lumbar vertebra such as BV/TV ratio, the cortical shell thickness, yield strength of bone, geometric parameters, and so on. These variables should be treated as the random variables, and in addition, as it was already mentioned, these random variable are dependent. Therefore, at least the correlation coefficients or covariances between these random variables should be known. However, we could not find any information about the probability distribution laws of the mentioned variables that affect lumbar vertebra fracture. Therefore, the Gaussian laws were applied to the modelling.

The proposed model can be used to determine the fracture risk of individual patients by applying the peculiar anatomical properties of lumbar vertebrae, such as TBS or BV/TV. So, the proposed evaluation scheme of assessing fracture risk includes the model of a vertebra with cancellous bone and the cortical shell, in-silico numerical results obtained by using FEM software, and the basic expressions of evaluating the fracture risk of a mechanical system. In case of the existence of other universally recognized methods, our proposed method can be used as a supplementary method with other known fracture risk evaluation methods.

3.4. Influence of Vertebra Cortical Shell Buckling

Mechanical properties. Different properties are assigned for particular phases. The dense cortical shell phase is characterised by volume density $\rho_{cor} = 1850 \text{ kg/m}^3$ [37,38], while the density of fully degenerated cancellous bone is $\rho_{can} = 100 \text{ kg/m}^3$. The cortical phase is modelled as an isotropic elastic-plastic continuum. The elastic properties are defined by elasticity modulus $E_{cor} = 8.0 \text{ GPa}$, and the Poisson's ratio is $\nu_{cor} = 0.3$ [14,39]. The plastic properties are defined on the basis of perfect plasticity and obey the von Mises yield criteria. The yield stress is $\sigma_Y = 64 \text{ MPa}$ [14]. This value is further used as the ultimate strength constant. The trabecular phase is modelled as an elastic orthotropic continuum. The elastic modulus of the trabecular bone in the vertical (longitudinal) direction is calculated according to the formula given in Reference [40] as follows:

$$E_{can,zz} = 4.73\rho_{can}^{1.56} \quad (13)$$

Thereby, the transverse elastic modulus is assumed to be the fraction of the longitudinal modulus, thus

$$E_{can,xx} = E_{can,yy} = 0.1E_{can,zz} \quad (14)$$

The Poisson's ratio is $\nu_{can,xy} = 0.3$, and $\nu_{can,xz} = \nu_{can,yz} = 0.2$.

Finite element method. The thin-walled domain of the cortical shell was discretised by shell finite elements. The shell element applied is a four-node element with six degrees of freedom at each node. Such an element is associated with plasticity and larger strain and describes structure buckling. It is suitable for analysing thin to moderately thick shell structures. The finite element mesh of cortical shells contains 3094 nodes and 2976 shell elements.

A cancellous bone, endplates, and posterior bone models were meshed with volumetric finite elements. This type of solid element is a higher-order 3-D 20-node solid element that allows quadratic displacement behaviour. The element supports plasticity, large deflection, and large strain capabilities. Finally, the solid phase was described by a 3-D mesh containing 348,138 nodes and 147,814 solid elements. The meshed model is presented in Figure 10a.

Shell and volumetric domains may be connected in a different manner, and two computational finite element models were finally generated. The first model reflects a healthy cortex state. A connection between the cortical wall and trabecular phase is implemented as a contact of two solids (Figure 10b). In the perfect case, the connection is modelled as a bonded contact, where no sliding or separation between faces or edges is allowed. For the shell-solid constraint option, an internal set of force-distributed constraints between nodes on the shell edges and nodes on the solid surface

is created. In the model, each shell node acts as the master node and associated solid nodes act as slave nodes.

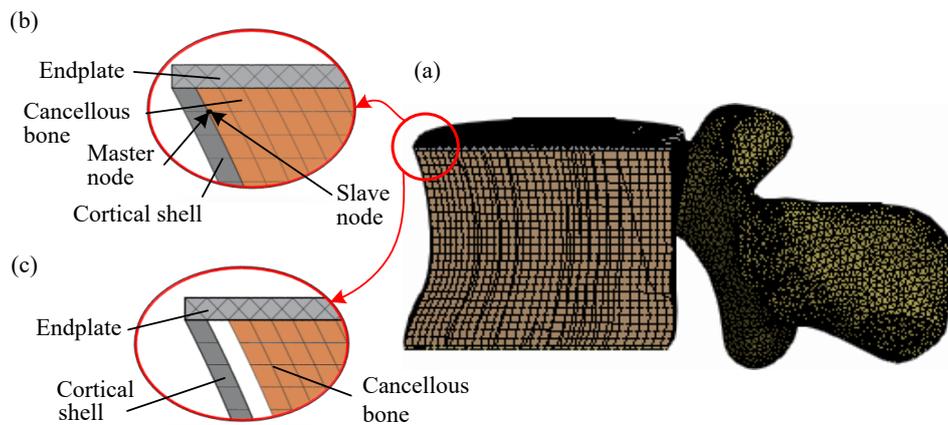


Figure 10. Cross-sectional view of the 3-D finite element model of the vertebra in the sagittal plane (a); bonded and unbonded, with a gap and cortical shells (b,c) respectively.

In the case of osteoporotic degradation, the bond is weakened. The second model reflects the limit case with a fully degenerated interface. The degradation effect could be evaluated by removing the connecting bond, and the gap was imposed between the shell and the solid (Figure 10c). Therewithin, integrity of the body was held by connecting the edge of the shell by endplates.

Numerical results. The physical nature of different models is qualitatively illustrated by deformed shapes, and a colour plot of the displacement magnitude of the cortical shell is shown in Figure 11. The displacement values, defined in millimetres, are illustrated in a unified colour scale. The first line of subfigures (Figure 11a–c) illustrates bonded shell-solid contact, while the next line of subfigures (Figure 11d–f) illustrates unbonded contact. The first column (Figure 11a,d) reflects results obtained with the large thickness equal to 0.5 mm, while the third column (Figure 11c,f) reflects results for most degradation of cortical shell. Characterising deformation shapes in a colour scale clearly illustrates the degradation degree. Unbonded contact leads to the occurrence of two higher-order deformation modes. Extremely high displacement, which exceeds nominal values by more than 1.5 times, is observed in the vicinity of point B. This result indicates buckling of the shell in the vicinity of this point.

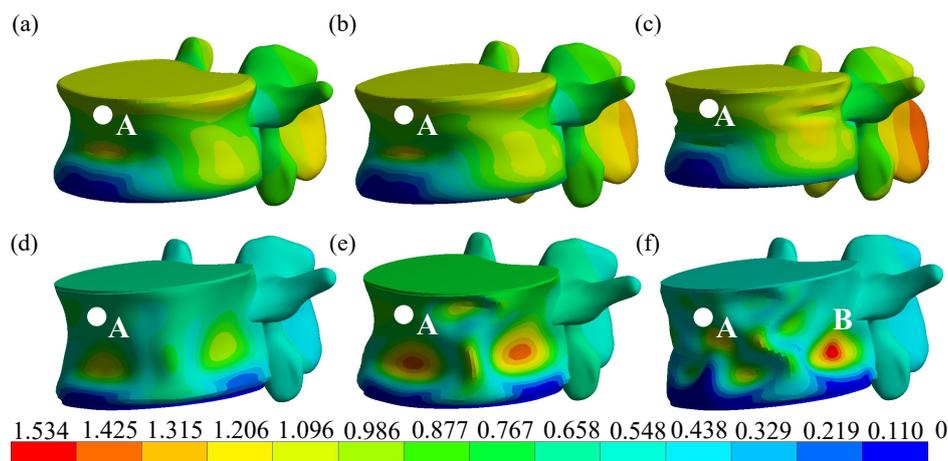


Figure 11. Deformed shape effect at the first bifurcation point A. Figure a–c for the bonded trabecular bone (Figure 10b): (a) when $\delta = 0.5$, $t = 0.86t_{max}$, (b) when $\delta = 0.4$ mm and $t = 0.85t_{max}$, (c) when $\delta = 0.2$ mm and $t = 0.58t_{max}$. Figure d–f for the unbounded trabecular bone (Figure 10c): (d) when $\delta = 0.5$ mm and $t = 0.32t_{max}$, (e) when $\delta = 0.4$ mm and $t = 0.32t_{max}$, (f) when $\delta = 0.2$ mm and $t = 0.38t_{max}$.

The most important results of the mechanical analyses are illustrated by considering the relative variation of vertebrae height h , %. According to medical practice [41], this parameter is used as a deformity grade. It is used as a fracture risk indicator. The normal state of vertebrae is characterised by deformability grade 0 which corresponds to variation $0 \leq h \leq 20\%$. It is worth noting that the obtained values of $h = 1.26 - 3\%$ are below the threshold value. The h threshold is 20%. Furthermore, the first bifurcation points of investigated vertebrae cortical shell is approximately twice higher than the maximum external loads applied in the proposed model (Figure 12). Therefore, it is obvious that buckling can be avoided in the current fracture risk model.

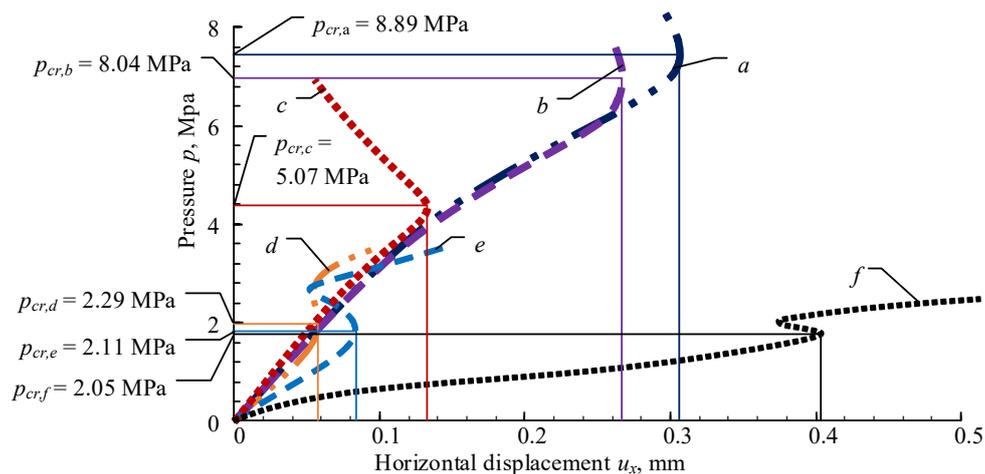


Figure 12. Load dependency graph. Displacement in horizontal direction at point A (Figure 11).

3.5. Discussion on Model Limitations

Creep. The deformations of bone under the load are time dependent not only due to osteoporosis but also due to creep [42–44]. Therefore, the long-term critical force or long-term strength of the bone is less than the short-term critical force or short-term strength. In case of the long-term deformation, long term deformation of the bone subject to long-term loading can be modelled properly using well-developed viscoelasticity or creep theories developed in the mechanics of materials provided that the characteristics of the bone is known, for example, modulus of elasticity, compliance function, or creep function. However, in the case of osteoporosis, the common viscoelasticity theories are not suitable due to fact that, in the classic viscoelasticity theories, the derivatives of the modulus of elasticity and the creep functions with respect to the time are nonnegative. In other words, the modulus of elasticity is a non-decreasing function with respect to time. However, in the case of osteoporosis, these conditions are violated and the long term analysis of the bone deformations when the osteoporosis simultaneously takes place is a much more complicated problem that also demands profound theoretical reconsiderations of the existing creep or viscoelasticity theories. The raised issue demands separate profound investigations which is not possible in the present study.

Anisotropy. Trabeculae bone tissues are strongly anisotropic materials with individual mechanical and topological properties for each patient. Good known representative volume elements (RVE) and homogenization-based methods [45] require topology of trabeculae, which needs micro-CT or micro-MRI scans. In the case of lumbar vertebra, they are impossible or very expensive. Therefore, modelling of mechanical behaviour of patient-specific lumbar vertebra is the challenge and we need some indirect methods for evaluating mechanical properties of patient-specific lumbar vertebra trabeculae. Wolf's law [46] helps to understand the anisotropy of mechanical properties of bone, and from theoretical point of view, we can still investigate simplified trabecular structures arranged in the direction of external stresses as a suitable mechanical model if we want to understand mechanical behaviour of trabeculae inside a vertebral body. On the other hand, it is experimentally observed fact

that, in vertebra L1, trabeculae arrangement is remodelled to withstand maximal compression loads. So, choosing of numerical compression tests is logical too.

Yield criteria. In general, lumbar vertebrae are strongly anisotropic plastic biological tissue. Therefore, failure criteria should be based on yield surface, which can be expressed as a Taylor expansion [47]

$$\left(\sum_{ik} a_{ik} \sigma_{ik} \right)^{\alpha} + \left(\sum_{pqmn} a_{pqmn} \sigma_{pq} \sigma_{mn} \right)^{\beta} + \left(\sum_{rstlmn} a_{rstlmn} \sigma_{rs} \sigma_{tl} \sigma_{mn} \right)^{\gamma} \leq 1 \quad (15)$$

where $a_{ik}, a_{pqmn}, a_{rstlmn}$ are strength tensors of different order. von Mises [48] proposed yield criterion for isotropic materials. Hill [49] extended the von Mises yield criterion of isotropic materials to anisotropic materials. Later Tsai [50] extended this criterion for anisotropic materials to a unidirectional lamina. It is easy to show that von Mises or Tsai–Hill failure criteria can be derived from Goldenblat–Kopnov failure criteria. Two-dimensional failure propagation simulations by Korenczuk et al. [51] showed that the von Mises failure criterion did not capture the failure type, location, or propagation direction nearly as well as the Tsai–Hill criterion. Both the von Mises and Tsai–Hill failure criteria severely underpredicted the amount of displacement needed to produce initial failure in the porcine abdominal aortas. Therefore, using of von Mises or Tsai–Hill failure criteria would overestimate lumbar vertebra fracture risk.

Remodelling. One-time resonant destruction of vertebra is a very rare clinical dysfunction, which was not investigated in our model. On the other hand, according to Wolf's law, long range loads enable the increase of bone strength in the greatest load direction and opposite decrease strength in lowest load direction. Therefore, we would take into account additionally bone remodelling or its fatigue property, but this article does not include an investigation of vertebra remodelling.

3.6. Validation of Mechanical Model

Extracting physical and micro-geometrical properties of vertebrae is very complicated. If we want to know micro-geometry of human bones, quantitative computational tomography (QCT) can be used on leg or hand bones but not on whole human spine or one of vertebra. Therefore, an alternating methodology of investigation of vertebrae mechanical properties can be 3-D printing. As is stated in Reference [52], today, 3-D printing is practical, useful technology in surgical planing or spine implantation. Obviously, knowing mechanical characteristics, 3-D printed vertebrae is the basic step in successful applications of such kind of structures.

The validation of numerical modelling was aided in our case by a physical test of the printed polylactide (PLA) vertebrae geometrical model (Figure 13a). The whole procedure took the following steps:

1. The printing of the vertebrae model;
2. The printing of the cylindrical PLA sample for mechanical properties of PLA to define;
3. The compression test of the printed PLA sample;
4. The compression test of the printed PLA vertebrae;
5. The determination of the mechanical properties of PLA by verifying the obtained load-displacement curve of the cylindrical sample;
6. The implementation of the defined stress-strain curve of PLA onto the numerical calculation;
7. The comparison of the results of the experimental and numerical studies.

The printing process was realized by using self-made 3-D printer, based on Prusa-i3 model (Figure 13b). The selection of this device was justified by its parameters: printer type FDM, build volume, $25 \times 21 \times 20 \text{ cm}^3$, minimal layer height $50 \text{ }\mu\text{m}$, and nozzle diameter 0.2 mm . The printing process took about 20 h, while the printing was completed. The properties of the osteoporotic sample of vertebrae was as follow: thickness of cortical shell, 0.5 mm , 0.1 BV/TV ratio, and 1.2 TBS .

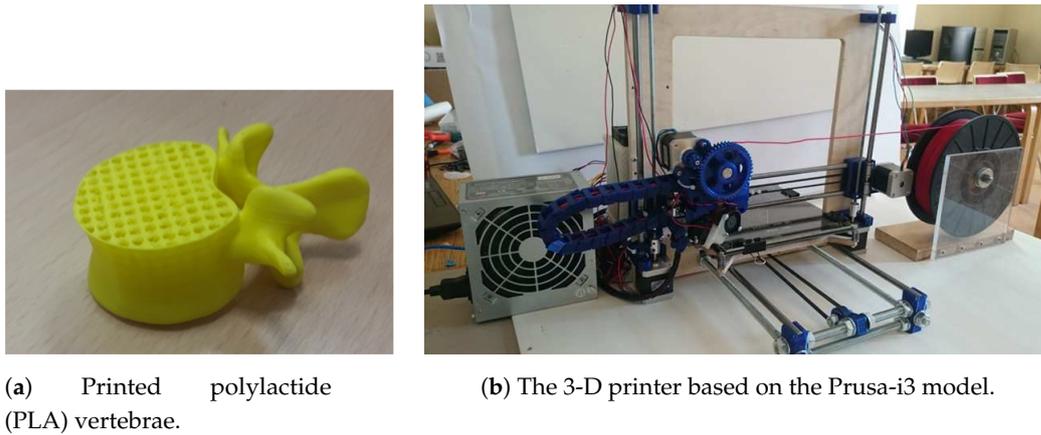


Figure 13. The 3-D printer and printed vertebra.

The physical experiment was performed by using a MultiTest 2.5-i compression machine (Figure 14). The selection of this device was based on suitable range of load 0–2500 N and relatively high precision ($\pm 0.1\%$ of full scale). The sample vertebra was tested under the constant loading speed (10 mm/min), and the displacement values were obtained: They are presented in Figure 15a. The same test was accomplished for the printed cylindrical sample too.

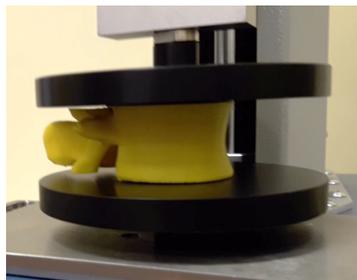
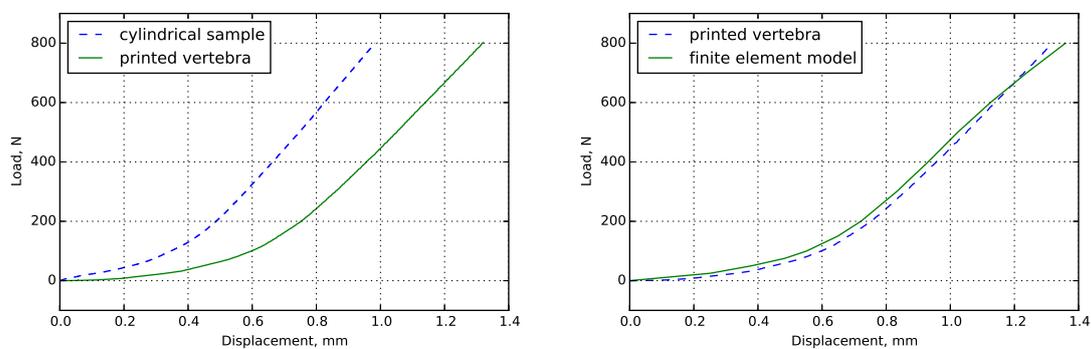


Figure 14. Testing of printed vertebrae by the compression machine.

As Figure 15a shows, the nonlinear behaviour of the printed PLA was revealed. The mechanical properties of PLA were determined by processing the load-displacement curve of the cylindrical sample (20 mm length, 15 mm diameter), and the stress–strain relation of PLA was defined.

The defined stress–strain curve of PLA (Figure 15a) was integrated into the numerical model of the vertebrae, and the finite element study was compared to the results of the physical experiment (Figure 15b).



(a) Compression test of cylindrical PLA sample and (b) Compression test of 3-D printed and FEM models of lumbar vertebrae.

Figure 15. The load-displacement curves.

As Figure 15b shows, the curves present corresponding characters, and the biggest difference between the results is constant (about 4%) until the value of load becomes 700 N. Then, the curves demonstrate similar magnitudes of displacement values, which does not exceed 5%. These results approve the adequacy of finite element model and shows that the method offered for osteoporotic bone validation is reliable.

4. Conclusions

The proposed method based on the cancellous bone and cortical shell in silico finite element modelling includes basic principles of evaluating fracture risk of the mechanical system and can be used after clinical researches to estimate the quantiles of the load-bearing capacity of a lumbar vertebra as well as the fracture risk of individual patients by applying the peculiar anatomical properties of the lumbar vertebra.

The results show that the fracture risk is substantially higher at relatively low levels of apparent BV/TV ratios and critical due to thinner cortical shells, suggesting high risk levels even during daily activities of typical external loads.

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Abbreviations

The following abbreviations are used in this manuscript:

FEM	Finite element method
FE	Finite element
IVD	intervetebral disk
BV/TV	Bone volume vs total volume
QCT	Quantitative computational tomography
BMD	Bone mineral density
DXA	Dual-energy X-ray absorptiometry

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